



## INTERNATIONAL JOURNAL FOR RESEARCH

IN APPLIED SCIENCE & ENGINEERING TECHNOLOGY

Volume: 7 Issue: IX Month of publication: September 2019

DOI: http://doi.org/10.22214/ijraset.2019.9127

www.ijraset.com

Call: © 08813907089 E-mail ID: ijraset@gmail.com

ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177

Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

# Comparative Autocorrelation Study for Foreign Exchange Reserve and Merchandise Export in India: Development of Prediction Model

Avik Ghosh

CMA, B.E. (Electrical), JAIIB, CAIIB, Diploma in Investment Law

Abstract: The research is having an objective to analyse the monthly data of Foreign Exchange Reserve of India and merchandise export from India since 1990-91. The analysis was aimed at finding an autocorrelation of the said variables. Both the variables are considerably significant in the broader domain of macroeconomic framework. Foreign exchange reserve classifies the strength of government reserve in other currencies to mitigate unforeseen circumstances whereas higher export value signifies better current account and trade position.

These parameters have strong correlation with the exchange rate of home currency and on the balance of payment. The degree of autocorrelation exemplifies the dependence of these variables on their past outcome.

This study tested the autocorrelation of the variables successfully and through the detailed statistical methodology of unit root testing, ARMA modelling and GARCH modelling, it established a statistical model that emanates the possibility to forecast the variables.

This outcome may have significant impact in policy making and predicting future trade trend.

Keywords: Foreign Exchange Reserve in India, Export from India, Autocorrelation, Unit Root Test, ARMA model, GARCH model, EGARCH model

### I. INTRODUCTION

The data on foreign exchange reserve position of a country keeps updating us on multiple facets of macroeconomy- the strength of home currency, the macroeconomic risk-taking ability, capability to fund import, trade contingency position and so on. This is even more significant for import-dependent country like India.

The data on merchandise export of any country emphasises its global trade position, trade surplus / deficit tendency, current account position, exchange rate position etc. When trade deficit exists in a country, it subsequently impacts current account deficit which is extremely relevant in modern market economy.

The composition of foreign exchange reserve (Table 1) depicts that the maximum weightage is for the foreign currency asset held by government. When the current account deficit cannot be compensated by FDI or FPI inflows, the foreign exchange reserves play a key role for macroeconomic balance.

The foreign exchange reserve also indicates the availability of contingency funds with the government to provide import support in case of expected economic stress.

The FDI and FPI can't always compensate the gap of current account deficit resulting in requirement of external borrowing or funding from own reserves. The 93%-95% contribution of foreign currency assets in total foreign exchange reserve explains the ongoing valuation of foreign currency in comparison with home currency and its changes reflect the situational improvement / degradation of home currency.

This has some correlation with the export amount from the country due to its capability to strengthen home currency. Table 1 also depicts an approximate 200% increase in foreign currency reserve of India in April 2019 in comparison with April 2005. Furthermore, the two variables, namely export value and foreign exchange reserve, have been considered for review and analysis to invigorate the meta-data analysis for finding out inner traits of the datasets for forecasting.



ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177 Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

Vaar	Month	Foreign Currency Assets		Go	Gold		Reserve Tranche Position		SDRs	
Year		(USD	% of	(USD	% of	(USD	% of	(USD	% of	(USD
		Million)	Reserve	Million)	Reserve	Million)	Reserve	Million)	Reserve	Million)
2019	Apr	390966	93.4%	23022	5.5%	3351	0.8%	1454	0.3%	418793
2018	Apr	395277	94.0%	21662	5.2%	2056	0.5%	1523	0.4%	420517
2017	Apr	349056	93.5%	20439	5.5%	2347	0.6%	1460	0.4%	373302
2016	Apr	339025	93.4%	20043	5.5%	2471	0.7%	1511	0.4%	363049
2015	Apr	327153	93.0%	19336	5.5%	1317	0.4%	4063	1.2%	351869
2014	Apr	283707	91.2%	20966	6.7%	1838	0.6%	4475	1.4%	310986
2013	Apr	263322	89.6%	23974	8.2%	2240	0.8%	4356	1.5%	293892
2012	Apr	260839	88.5%	26618	9.0%	2915	1.0%	4474	1.5%	294846
2011	Apr	282037	90.0%	23790	7.6%	3013	1.0%	4671	1.5%	313511
2010	Apr	254773	91.1%	18537	6.6%	1341	0.5%	4982	1.8%	279633
2009	Apr	241487	95.9%	9231	3.7%	983	0.4%	1	0.0%	251702
2008	Apr	304225	96.8%	9427	3.0%	485	0.2%	18	0.0%	314155
2007	Apr	196899	96.3%	7036	3.4%	463	0.2%	11	0.0%	204409
2006	Apr	153598	95.6%	6301	3.9%	772	0.5%	6	0.0%	160677
2005	Apr	135950	95.8%	4443	3.1%	1443	1.0%	5	0.0%	141841

Table 1

### II. PREVIOUS RESEARCH

Various research work has been undertaken and accomplished earlier on the foreign exchange reserve and export value of multiple countries, but the autocorrelation study and modeling of the variables are hardly available in research domain. Bhattacharya B, Mookherjee J, in 2001 and Doong, S.-Ch., Yang, Sh.-Y., Wang, A., in 2005, in their research papers elaborated the aspects of the importance of foreign exchange reserve and its impact on trade deficit, exchange rate and even in stock market for emerging countries. Mohammad, S. D., Hussain, A., & Ali, A., in 2009 and Aizenman, J. and Marion, N., in 2002, assessed the impact of foreign exchange reserve on domestic economy in case of middle-east countries.

Disyatat, P., in 2001 and Greenspan, A, in 1991 stated the importance of foreign exchange reserve in macroeconomic stability. They also highlighted the examples of some emerging countries who were capable of maintaining strong reserve position and converted themselves to trade power house. They suggested to incorporate the provision of purchasing foreign currency by government to make the reserve position stronger and also advised that well-informed calculated purchase of the foreign currency is an excellent mode of investment and hedging.

Kenen, P. and Yudin, E., in 1965, and Ford, J.L. and Huang, G., in 1994 worked out the evolution of the concept of higher forex reserve by global superpowers. This got reclarified and more specific with the advent of globalization and liberalization. All these studies covered the aspects of forex reserve and its gradual change for many countries. These papers also justified the changing trend in the field of merchandise export and investment in foreign currency. This paper mostly covers the quantitative aspects of the meta-data to analyze the possibility of finding an autocorrelation of the variables.

### III. INITIAL THEORETICAL FRAMEWORK AND METHODOLOGY

This research work, as stated earlier, aimed at establishing relationship of foreign exchange reserve (FX) of India and export from India (EX) with their past values. In other words autocorrelation of the said parameters was tested in this paper. As the theoretical framework goes, the data points were collected from relevant sources and the analysis has been performed based on monthly data from FY 1990-1991 to the last available data (June / July 2019). The log returns for both the variables have been calculated for all analytical purposes due to higher chance of normalization. The log returns of Foreign Exchange Reserve (RFX) and Export from India (REX) data were initially placed in histograms to have an overview of normalization effect. Subsequently, the data series statistical exercises were performed that started with correlogram to analyze the existence of autocorrelation. This was revalidated by Breusch-Godfrey Serial Correlation LM test. The Unit Root Testing was performed next with Augmented Dicky Fuller test to verify the stationarity of the variables barring which the unit roots take the future prediction and data model equation away from the

ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177

Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

expected trajectory. Once the stationarity was established, both RFX and REX was modeled in terms of Autoregressive Moving Average (ARMA) models and the variables were both tested on ARMA(1,1) and ARMA(2,2) models. Once the acceptability of the models was decided on the probability factors and t-statistics at 5% significance level, it is concluded that the modeling of the variable with that specific autocorrelated model is possible. These models were further tested for their respective variances with General Autoregressive Conditional Heteroskedasticity (GARCH) and Exponential GARCH (EGARCH) models.

### IV. PRESENTATION OF DATA AND EXPLANATION

The initial representation of the data vide histograms (Figure 1) confirms that the normalization trend in the data set is present. The export data set is less skewed than foreign exchange data which is reconfirmed with higher value of Kurtosis and very high value of Jarque-Bera. Once the initial representation of the variables is performed, the autocorrelation tendency is tested with correlogram with 18 lags and the same was revalidated by Breusch-Godfrey Serial Correlation LM test. In both the tests, the hypothesis Null Hypothesis H0: There is no autocorrelation in the Indian forex reserve and export data and Alternate Hypothesis H1: There is autocorrelation in the Indian forex reserve and export data. The correlogram (Figure 2) for export data clearly indicates autocorrelation from first lag

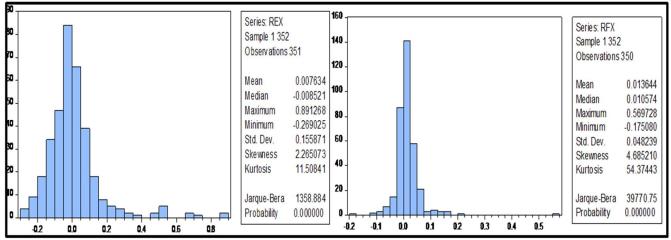


Figure 1

And the same is present from third lag in case of foreign exchange. This was established due to very low p-value and significant Q-statistics. Very low Chi-square value and significant F-statistic in Breusch-Godfrey Serial Correlation LM test (Figure 3) confirm the presence of autocorrelation in both the variables. This helps the analysis to be conducted further to derive the model for representing the variables for future prediction. The presence of Unit root in the data sets was checked with Augmented Dicky Fuller test. Unit roots disrupt stationarity of any variable and hinder the process of modelling the

Sample: 1 352 Included observation	ns: 351	•				Sample: 1 352 Included observation	ns: 350				
Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob	Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
		7 0.021 8 -0.150 9 0.032 10 -0.007 11 -0.290 12 0.680 13 -0.236 14 -0.011	-0.156 -0.054 -0.173 -0.113 -0.069 -0.020 -0.213 -0.156 -0.131 -0.503 0.435 0.087 -0.005	42.842 42.912 49.124 49.282 49.289 49.447 57.594 57.971 57.987 88.619 257.74 278.09 278.14 278.48	0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000			2 0.070 3 0.233 4 -0.026 5 0.041 6 -0.100 7 -0.062 8 -0.013 9 0.073 10 -0.123 11 0.049 12 0.080 13 -0.038	0.070 0.237 -0.020 0.006 -0.160 -0.067 -0.012 0.160 -0.094 0.026 0.003 0.066 -0.020	21.135 21.374 21.973 25.530 26.894 26.956 28.885 34.332 35.200 37.529 38.051 42.341 42.673	0.396 0.000 0.000 0.001 0.000 0.001 0.000 0.000 0.000 0.000 0.000
1 p1 1 p1		17 0.048 18 0.057		287.72 288.94		- 1		17 0.138 18 0.012		49.928 49.978	

Figure 2



ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177 Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

Variable for forecasting. The tests were performed for both the variables with only constant and constant with linear trend. The results (Figure 4) highlight that t-statistic (calculated t-value) or  $t_{stat} < t_{critical}$  obtained from the ADF table. This rejects the null hypothesis and confirms the alternate hypothesis of no unit root for both the variables. As the stationarity of the variables are established due to absence of unit roots, the variables were experimented to be modelled with ARMA model. Both the

Breusch-Godfrey Seri Null hypothesis: No s			Breusch-Godfrey Serial Correlation LM Test Null hypothesis: No serial correlation at up to 6 lags					
F-statistic Obs*R-squared	7.052116 38.43487	Prob. F(6,343) Prob. Chi-Square(6)		F-statistic Obs*R-squared		Prob. F(6,344) Prob. Chi-Square(6)	0.0000	

Figure 3

Variables were tested with ARMA(1,1) (Figure 5) and ARMA(2,2) (Figure 6) models and the results, with the help of probability and t-statistics, depicted that export data can be modelled with ARMA(1,1) whereas foreign exchange reserves data can be represented as ARMA(2,2) model. The derived co-efficient in both the cases resulted in the representation of the variables with its past values due to the

Null Hypothesis: REX ha Exogenous: Constant Lag Length: 11 (Automat		axlag=16)	Null Hypothesis: RFXhas a unit root Exogenous: Constant Lag Length: 9 (Automatic - based on SIC, maxlag=16)						
		t-Statistic	Prob.*			t-Statistic	Prob.*		
Augmented Dickey-Fulle	r test statistic	-7.607668	0.0000	Augmented Dickey-Ful	ler test statistic	-5.817940	0.0000		
Test critical values: 1% level		-3.449447		Test critical values:	1% level	-3.449389			
	5% level	-2.869850			5% level	-2.869825			
	10% level	-2.571266			10% level	-2.571253			
*MacKinnon (1996) one-	sided p-values.			*MacKinnon (1996) one-sided p-values.					
Null Hypothesis: REX ha	is a unit root			Null Hypothesis: RFX has a unit root Exogenous: Constant, Linear Trend Lag Length: 9 (Automatic - based on SIC, maxlag=16)					
Exogenous: Constant, Li									
Lag Length: 11 (Automat		axlag=16)							
		t-Statistic	Prob.*			t-Statistic	Prob.*		
Augmented Dickey-Fulle	r test statistic	-7.607006	0.0000	Augmented Dickey-Ful	ler test statistic	-6.446699	0.0000		
Test critical values:	1% level	-3.985361		Test critical values:	1% level	-3.985280			
	5% level	-3,423136			5% level	-3.423097			
	10% level	-3.134497			10% level	-3.134474			
*MacKinnon (1996) one-	sided p-values.		*MacKinnon (1996) one-sided p-values.						

Figure 4

Presence of autocorrelation. The export data and the foreign exchange reserve data can be represented as:  $REX_t = 0.008328 + \epsilon_t + 0.390452REX_{t-1} - 0.860746\epsilon_{t-1}$  in case of export data set at any time period t and  $RFX_t = 0.013451 + \epsilon_t + 1.271108RNE_{t-1} - 0.787041RNE_{t-2} - 1.312159\epsilon_{t-1} + 0.949649$   $\epsilon_{t-2}$  in case of

Dependent Variable: Rt Method: ARMA Maxim un Sample: 2 352 Included observations: Convergence achieved Coefficient covariance of	n Likelihood (C 351 after 39 iteratio	ons	Dependent Variable: RFX Method: ARMA Maximum Likelihood (OPG - BHHH) Sample: 2 351 Included observations: 350 Convergence achieved after 34 iterations Coefficient covariance computed using outer product of gradients						
Variable	Coefficient	Std. Error	t-Statistic	Prob.	Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(1) MA(1) SIGMASQ	0.008328 0.390452 -0.860746 0.019222	0.002056 0.065874 0.048759 0.001176	4.049877 5.927248 -17.65291 16.35041		C AR(1) MA(1) SIGMASQ	0.013644 -0.160254 0.137368 0.002319	0.003293 1.756745 1.763103 7.97E-05	4.142704 -0.091222 0.077912 29.11297	0.0000 0.9274 0.9379 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.206594 0.199735 0.139439 6.746757 195.1318 30.11832 0.000000	5 S.D. dependent var 0.15 9 Akaike info criterion -1.08 8 Hannan-Quinn criter1.07 2 Durbin-Watson stat 2.05		0.007634 0.155871 -1.089070 -1.045073 -1.071560 2.057968		0.000540 -0.008125 0.048435 0.811689 565.0212 0.062360 0.979613	Mean depende S.D. depende Akaike info cr Schwarz crite Hannan-Quin Durbin-Watso	ent var iterion rion in criter.	0.013644 0.048239 -3.205835 -3.161745 -3.188286 1.989288
Inverted AR Roots Inverted MA Roots	.39 .86				Inverted AR Roots Inverted MA Roots	16 14			

Figure 5



ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177 Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

Foreign exchange reserve data set. These equations may also be used as forecasting tool for the said variables. Due to the acceptability of ARMA-fit models in case of both the variables, the variances can

Dependent Variable: REX Method: ARMA Maxim um Likelihood (OPG - BHHH)

Sample: 2 352

Included observations: 351

Convergence achieved after 44 iterations

Coefficient covariance computed using outer product of gradients Variable Coefficient Std. Error t-Statistic Prob. 3.627813 C 0.008359 0.002304 0.0003 AR(1) 0.288949 0.763234 0.378585 0.7052 AR(2) 0.132244 0.296508 0.446004 0.6559 MA(1) -0.791899 0.768715 -1.030159 0.3037 MA(2) -0.075933 0.658386 -0.115333 0.9082 SIGMASQ 0.001154 0.019064 16.52463 0.0000 R-squared 0.213077 Mean dependent var 0.007634 Adjusted R-squared 0.201672 0.155871 S.D. dependent var -1.085989 S.E. of regression 0.139270 Akaike info criterion Sum squared resid 6.691629 Schwarz criterion -1.019993196.5911 Hannan-Quinn criter. -1.059723Log likelihood F-statistic 18 68330 Durbin-Watson stat 1.990152 Prob(F-statistic) 0.000000 Inverted AR Roots .54 -.25 .88 -.09 Inverted MA Roots

Dependent Variable: RFX

Method: ARMA Maximum Likelihood (OPG - BHHH)

Sample: 2 351

Included observations: 350

Convergence achieved after 165 iterations

Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	0.013451	0.004015	3.349907	0.0009
AR(1)	1.271108	0.050129	25.35667	0.0000
AR(2)	-0.787041	0.044259	-17.78247	0.0000
MA(1)	-1.312159	0.040670	-32.26320	0.0000
MA(2)	0.949649	0.036369	26.11116	0.0000
SIGMASQ	0.002129	7.33E-05	29.03352	0.0000
R-squared	0.082438	Mean depend	lent var	0.013644
Adjusted R-squared	0.069102	S.D. depende	0.048239	
S.E. of regression	0.046543	Akaike info cri	-3.275955	
Sum squared resid	0.745177	Schwarz criter	-3.209819	
Log likelihood	579.2922	Hannan-Quin	n criter.	-3.249631
F-statistic	6.181339	Durbin-Watso	n stat	2.065862
Prob(F-statistic)	0.000017			
Inverted AR Roots	.6462i	.64+.62i	•	
Inverted MA Roots	.66+.72i	.6672i		

Figure 6

Dependent Variable: REX

Method: ML ARCH - Normal distribution (OPG - BHHH / Marquardt steps)

Sample (adjusted): 3 352

Included observations: 350 after adjustments

Failure to improve likelihood (non-zero gradients) after 78 iterations

Coefficient covariance computed using outer product of gradients

MA Backcast 2

Inverted MA Roots

Presample variance: backcast (parameter = 0.7) GARCH = C(4) + C(5)\*RESID(-1)\*2 + C(6)\*GARCH(-1)

84

Coefficient z-Statistic Prob. Variable Std. Error C 0.007693 0.004833 1 591949 0 1114 AR(1) 0.203921 0.318836 1.563527 0.1179MA(1) -0.843408 0.111375 -7.572678 0.0000 Variance Equation C 0.014533 0.015052 0.965526 0.3343 RESID(-1)^2 0.148781 0.147271 1.010249 0.3124 GARCH(-1) 0.598781 0.374397 1.599320 0.1097 0.202924 0.008052 R-squared Mean dependent var Adjusted R-squared 0.198330 0.155897 S.D. dependent var 0.139584 S.E. of regression Akaike info criterion -0.864041 Sum squared resid 6.760873 Schwarz criterion -0.797905 Log likelihood 157.2071 Hannan-Quinn criter -0.837716 Durbin-Watson stat 1.931137 Inverted AR Roots .32

Dependent Variable: REX Method: ML ARCH - Normal distribution (OPG - BHHH / Marquardt steps) Sample (adjusted): 3 352 Included observations: 350 after adjustments

Convergence achieved after 45 iterations

Coefficient covariance computed using outer product of gradients

MA Backcast: 2

Presample variance: backcast (parameter = 0.7)

LOG(GARCH) = C(4) + C(5)\*ABS(RESID(-1)/@SQRT(GARCH(-1))) + C(6)\*RESID(-1)/@SQRT(GARCH(-1)) + C(7)\*LOG(GARCH(-1))

Variable Coefficient Std. Error z-Statistic Prob. C 0.5603 -0.001888 0.003242 -0.582416 AR(1) 0.636126 0.069953 9.093615 0.0000 MA(1) -0.868304 0.046232 -18.78126 0.0000 Variance Equation C(4) -4.360645 -12.02755 0.0000 0.362555 0.811706 0.0000 C(5) 0.109094 7.440449 -0.502787 0.0000 0.089869 -5.594638 C(6) 0.144561 0.008052 R-squared 0.122144 Mean dependent var Adjusted R-squared 0.117085 S.D. dependent var 0.155897 0.146487 -1.553749 S.E. of regression Akaike info criterion 7.446053 -1.476590 Sum squared resid Schwarz criterion Log likelihood 278,9060 Hannan-Quinn criter. -1.523037 Durbin-Watson stat 2.393216 Inverted AR Roots .64 87 Inverted MA Roots

Figure 7

Also be modelled. We have tested the variance modelling for both the variables with GARCH and EGARCH models. For export dataset, the EGARCH model complies to the significance of the model at 5% level (Figure 7) for all coefficients except c (7) but the GARCH model fails to do so for all coefficients. With the feasibility of the EGARCH model the variance for export data points can be expressed as:

$$\log (\sigma_t^2) = -4.360645 + 0.144561\log (\sigma_{t-1}^2) - 0.502787 \ \epsilon_{t-1} / (\sqrt{\sigma_{t-1}^2}) + 0.811706 \ [ \ (I\epsilon_{t-1}I/\sigma_{t-1}^2) - \sqrt{\frac{2}{\pi}} \ ]$$



ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177 Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

Dependent Variable: RFX Method: ML ARCH - Normal distribution (OPG - BHHH / Marquardt steps) Method: ML ARCH - Normal distribution (OPG - BHHH / Marquardt steps) Sample (adjusted): 4 351 Sample (adjusted): 4 351 Included observations: 348 after adjustments Included observations: 348 after adjustments Failure to improve likelihood (non-zero gradients) after 53 iterations Convergence not achieved after 500 iterations Coefficient covariance computed using outer product of gradients Coefficient covariance computed using outer product of gradients MA Backcast: 23 Presample variance: backcast (parameter = 0.7) Presample variance: backcast (parameter = 0.7) LOG(GARCH) = C(6) + C(7)\*ABS(RESID(-1)/@SQRT(GARCH(-1))) + C(8)  $GARCH = C(6) + C(7)*RESID(-1)^2 + C(8)*GARCH(-1)$ \*RESID(-1)/@SQRT(GARCH(-1)) + C(9)\*LOG(GARCH(-1)) Variable Coefficient Std. Error z-Statistic Prob. Variable Coefficient Std. Error z-Statistic Prob. C 0.015155 0.001788 8.477323 0.0000 C 0.005425 0.001525 3.556421 0.0004 AR(1) 0.425863 0.140294 3.035495 0.0024 -2.980424 0.331681 AR(2) 0.097530 3.400804 0.0007 AR(2) 0.608841 0.083834 7.262468 0.0000 MA(1) -0.017126 0.139843 -0.122467 0.9025 0.0000 MA(2) 0.0000 -0.2891410.054302 -5.324677MA(2) -0.3578640.096336 -3.7147480.0002 Variance Equation Variance Equation 0.086728 0.013848 6.262933 0.0000 4.71E-05 C(6) 3.98E-06 11.83516 0.0000 C(7) 0.006803 0.020068 0.339016 0.7346 RESID(-1)<sup>2</sup> GARCH(-1) 0.014425 10.29639 21.76018 0.148530 0.0000 C(8) C(9) 0.598530 0.027506 -0.071301 0.012364 -5.766722 0.0000 0.0000 1.009751 6.33E-07 1594097 0.0000 R-squared -0.147171Mean dependent var 0.013648 -0.073207 0.013648 R-squared Mean dependent vai Adjusted R-squared -0.160549 S.D. dependent var 0.048376 Adjusted R-squared -0.085722S.D. dependent var 0.048376 S.E. of regression 0.052115 Akaike info criterior -3.566153 0.050407 Akaike info criterion -4.288179 S.E. of regression Sum squared resid 0.931574 Schwarz criterion -3.477596 Sum squared resid 0.871511 Schwarz criterion -4.188553 628.5106 Log likelihood Hannan-Quinn criter. -3.530897og likelihood Hannan-Quinn criter. -4.248516 Durbin-Watson stat 2.743922 Durbin-Watson stat 2.522480 Inverted AR Roots .83 -.40 Inverted AR Roots Inverted MA Roots

Figure 8

This equation is useful to evaluate future variances as well. While assessing similar data points for foreign exchange reserve in India, the variances can be well-expressed by GARCH and EGARCH models. However, GARCH model fits it with all coefficients but EGARCH has restriction with c (7). Hence the variances have been considered to be modelled with GARCH model and can be represented as:  $\sigma_t^2 = 4.71e^{-5} + 0.148530 \, \epsilon_{t-1}^2 + 0.598530 \, \sigma_{t-1}^2$ 

### A. Source of Data

The above analysis was based on data available at Reserve Bank of India official data portal (dbie.rbi.org.in). The selected data set contains forex reserve data and export data since 1990-1991 and has more than 350 data points each. The subsequent analysis has been performed with EViews software.

### B. Scope and Limitations of Research

The research work covers two key aspects of macroeconomic indicators and trade balance health of any country. In case of India, the study has been performed to assess the dependency of present data on its past. The same analysis could have been performed for few other countries to ascertain a comparative positioning. Few other parameters related with the present research work namely import, current account deficit etc. have not been included to maintain the objective orientation. More data points could have been useful to derive the coefficients more precisely, but could not be done due to unavailability of the data beyond 1990-91.

### V. CONCLUSION

This research work is a combination of identification of macroeconomic indicators, analysis of its relevance with its past data through statistical tools and modelling its forecasting equation with econometrics application. The study established that foreign exchange reserve and merchandise export values for India are autocorrelated parameters. Hence, for both the parameters, past value means a lot to predict future values resulting in an authentic and statistically modelled guideline for forecasting. This autocorrelation for forex reserve emanated that the present reserve position would impact the future reserve position by reducing the probability of volatility impact to a great extent. Similarly, the export data analysis and the autocorrelation established from that proves the importance of improving the present position to have a reasonably balanced future. The models and equations established are extremely helpful for future projection that may be used as a guiding tool. This assessment may also help the policy makers to change their focus on policy intervention and initiatives resulting in lesser dependence on market volatility and more focussed approach in predictive model-dependent policy measures to have an improved trade balance.



ISSN: 2321-9653; IC Value: 45.98; SJ Impact Factor: 7.177 Volume 7 Issue IX, Sep 2019- Available at www.ijraset.com

### REFERENCES

- [1] Bhattacharya B, Mookherjee J (2001), Causal relationship between and exchange rate, foreign exchange reserves, value of trade balance and stock market: case study of India. Department of Economics, Jadavpur University, Kolkata, India.
- [2] Doong, S.-Ch., Yang, Sh.-Y., Wang, A., 2005. "The dynamic relationship and pricing of stocks and exchange rates: Empirical evidence from Asian emerging markets," Journal of American Academy of Business, Cambridge, Vol.7, No1, pp.118-23.
- [3] Mohammad, S. D., Hussain, A., & Ali, A. (2009), Impact of Macroeconomics Variables on Stock Prices Emperical Evidance in Case of KSE. European Journal of Scientific Research, vol.38 no. 1, pp96-103.
- [4] Aizenman, J. and Marion, N. "The High Demand for International Reserves in the Far East: What's Going On?" Working Paper, UC Santa Cruz and the NBER; Dartmouth College, 2002
- [5] Beaufort and Kapteyn. "Reserve Adequacy in Emerging Market Economies" Working Paper, International Monetary Fund, 2001
- [6] Disyatat, P. "Currency Crisis and Foreign Reserves: A Simple Model" Working Paper, International Monetary Fund, 2001
- [7] Greenspan, A. 'Currency Reserves and Debt'. Remakrs at the World Bank Conference on Trends in Reserves. April, 1999.
- [8] Kenen, P. and Yudin, E. "The Demand for International Reserves" The Review of Economics and Statistics, 47, 1965: 242-250
- [9] Ford, J.L. and Huang, G. "The Demand for International Reserves in China: An ECM Model with Domestic Monetary Disequilibrium" Economica, 67, 1994: 379-97









45.98



IMPACT FACTOR: 7.129



IMPACT FACTOR: 7.429



### INTERNATIONAL JOURNAL FOR RESEARCH

IN APPLIED SCIENCE & ENGINEERING TECHNOLOGY

Call: 08813907089 🕓 (24\*7 Support on Whatsapp)